

# MATHEMATICAL FOUNDATIONS FOR FINANCE

## Introduction to Probability and Statistics: Part 2

Philip H. Dybvig  
Washington University  
Saint Louis, Missouri

## Continuous random variables

Discrete random variables take on isolated values, while continuous random variables can take on all values in some interval or intervals. Which to use is a modelling choice often based on convenience; taking on a thousand different values may not be qualitatively different from taking on any value in a continuum.

Work with a continuous random variable  $x$ , we often work with the probability density function  $f(x)$  or the cumulative distribution function  $F(x)$ . The probability density function tells the density of the probability measure with respect to  $x$ ; we can say informally that  $f(x_0)\Delta x$  is the approximate probability of  $x$  being in an interval of length  $\Delta x$  at  $x_0$ .

The cumulative distribution function  $F(x_0)$  at  $x_0$  is the probability  $x$  will take on a value less than  $x_0$ , and it is related to the density function  $f(x)$  by  $F(x_0) = \int_{x=-\infty}^{x_0} f(x)dx$ . We can express the expectation of a function  $h$  of  $x$  in terms of  $F$  or  $f$  as

$$E[h(x)] = \int_{x=-\infty}^{\infty} h(x)f(x)dx = \int_{x=-\infty}^{\infty} h(x)dF(x).$$

## Properties of density and distributions functions

- If the probability density function  $f(x)$  exists everywhere, then
  - $(\forall x) f(x) \geq 0$
  - $\int_{x=-\infty}^{\infty} f(x) = 1$
  - $(\forall x_0) F(x_0) = \int_{x=-\infty}^{x_0} f(x)$
  - $f(x)$  is not defined for discrete random variables.
- The probability distribution function has the the properties:
  - $(\forall x_0 \leq x_1) F(x_1) \geq F(x_0)$
  - $\lim_{x \rightarrow -\infty} F(x) = 0$
  - $\lim_{x \rightarrow \infty} F(x) = 1$
  - $F(x)$  is defined for all random variables, discrete or continuous.

## Using the density to compute moments

We know that  $E[h(x)] = \int_{x=-\infty}^{\infty} h(x)f(x)dx$ . In particular, this allows us to compute moments like the mean ( $h(x) = x$ ) and the variance ( $h(x) = (x - \mu_x)^2$ ), and other statistics like the skewness and kurtosis that can be written as functions of moments. Specifically,

$$\mu_x = E[x] = \int_{x=-\infty}^{\infty} x f(x) dx.$$

$$\begin{aligned} \text{var}(x) &= E[x^2] - \mu_x^2 = \left( \int_{x=-\infty}^{\infty} x^2 f(x) dx \right) - \mu_x^2 \\ \text{skew}(x) &= \frac{E[(x - \mu_x)^3]}{\sigma_x^3} = \frac{\int_{x=-\infty}^{\infty} (x - \mu_x)^3 f(x) dx}{\text{var}(x)^{3/2}} \\ \text{kurt}(x) &= \frac{E[(x - \mu_x)^4]}{\sigma_x^4} = \frac{\int_{x=-\infty}^{\infty} (x - \mu_x)^4 f(x) dx}{\text{var}(x)^2} \end{aligned}$$

## In-class exercise: moments from a density function

Let  $\lambda > 0$  be a given constant. Suppose  $x$  has the exponential distribution of the form  $f(x) = k \exp(-\lambda x)$  for  $x > 0$  and  $f(x) = 0$  for  $x \leq 0$ .

- Compute the constant  $k$ .
- Compute  $E[x]$ .
- Compute  $\text{var}(x)$ .
- Compute the cumulative distribution function  $F(x)$ .

## Normal distribution

The normal distribution (traditional bell curve) with parameters  $\mu$  and  $\sigma$  has the density function

$$f(x) = \frac{1}{\sqrt{2\pi}\sigma} \exp\left(-\frac{(x - \mu)^2}{2\sigma^2}\right).$$

A random variable  $x$  with this density has mean  $\mu$ , standard deviation  $\sigma$ , and variance  $\sigma^2$ . Sometimes this density  $f(x)$  is written as  $n(x)$ , and the corresponding cumulative distribution function  $F(x)$  is written as  $N(x)$ . We can write  $N(x_0) = \int_{x=-\infty}^{x_0} n(x)dx$  (as always), but we do not know a closed form expression for  $N$ .

The normal distribution plays an important part in statistical theory, because if we have a large number  $T$  of independent and identically distributed (i.i.d.) draws from a random variable  $y$  with mean  $m$  and standard deviation  $s$ , then  $(\sum_{t=1}^T (y_t - m)) / (s\sqrt{T})$  is approximately distributed normally with mean  $\mu = 0$  and standard deviation  $\sigma = 1$ . This fact is the starting point of many statistical tests.

## Normal distribution: moment generating function, skewness and kurtosis

The moment generating function of a random variable  $x$  is given by

$$M(t) = E[e^{tX}].$$

$M(t)$  is called the moment generating function because the  $n^{\text{th}}$  moment of  $x$  around the origin is equal to the  $n^{\text{th}}$  derivative of  $M(t)$  evaluated at 0. The moment generating function is also useful in finance because it allows us to calculate the expected utility of an exponential utility function.<sup>1</sup>

For a normal random variable  $x$  with mean parameter  $\mu$  and standard deviation parameter  $\sigma$ , the moment generating function is  $e^{\mu t + \sigma^2 t^2 / 2}$ . This can be used to verify the mean  $\mu$ , variance  $\sigma^2$ , skewness 0, and kurtosis 3 of  $x$ .

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<sup>1</sup>Arguably, the characteristic function,  $E[e^{itx}]$  (sometimes defined as a constant times this), where  $i^2 = -1$ , is more useful in general because it exists for all distributions. However, the moment generating function is suitable for our purposes.

## In-class exercise: normal moment generating function

Let  $x$  be a random variable with moment generating function  $M(t) = e^{\mu t + \sigma^2 t^2 / 2}$ . Show that

- $x$  has mean  $\mu$
- $x$  has variance  $\sigma^2$

Consider the utility function  $u(w) = -e^{-Aw}$  (exponential utility with absolute risk aversion  $A$ ). Derive  $E[u(x)]$ .

Extra for experts: show that the skewness of  $x$  is 0.



## Joint distributions and joint normal density

If we have more than one continuous random variable  $x_1, \dots, x_n$ , we can use the joint density  $f(x_1, \dots, x_n)$  for all the random variables to write expectations of functions of all of them:

$$E[h(x_1, \dots, x_n)] = \int_{x_1=-\infty}^{\infty} \int_{x_2=-\infty}^{\infty} \dots \int_{x_n=-\infty}^{\infty} h(x_1, \dots, x_n) dx_1 dx_2 \dots dx_n.$$

One example of this is the bivariate joint normal distribution, which has density

$$f(x, y) = \frac{1}{2\pi\sigma_x\sigma_y\sqrt{1-\rho^2}} \exp\left(\frac{-1}{2(1-\rho^2)} \left(\frac{(x-\mu_x)^2}{\sigma_x^2} - \frac{2\rho(x-\mu_x)(y-\mu_y)}{\sigma_x\sigma_y} + \frac{(y-\mu_y)^2}{\sigma_y^2}\right)\right)$$

In this density, the parameters  $\mu_x$ ,  $\mu_y$ ,  $\sigma_x$ , and  $\sigma_y$  are what they seem to be. The parameter  $\rho$  is the correlation coefficient, which equals

$$\rho = \frac{\text{cov}(x, y)}{\sigma_x\sigma_y}$$

and takes on values between  $-1$  and  $1$  (for all random variables, not just jointly normal variables).

## Independence

Two random variables are independent if knowing about one random variable gives us no information about the other. This would be true (or at least a very good approximation) for repeated rolls of dice or flips of a coin, and it is a good approximation for returns in some security markets.

In terms of the distribution functions,  $x$  and  $y$  are independent if the joint probability density is multiplicatively separable:

$$p(x, y) = f(x)g(y)$$

In this case,  $f(x)$  is the density function of  $x$  and  $g(y)$  is the density function of  $y$ .

In-class exercise: independence and joint normality

Prove that  $x$  and  $y$  are independent if they are joint normally distributed with zero correlation ( $\rho = 0$ ).